

# Package ‘covdepGE’

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**Description** A covariate-dependent approach to Gaussian graphical modeling as described in Dasgupta et al. (2022). Employs a novel weighted pseudo-likelihood approach to model the conditional dependence structure of data as a continuous function of an extraneous covariate. The main function, `covdepGE::covdepGE()`, estimates a graphical representation of the conditional dependence structure via a block mean-field variational approximation, while several auxiliary functions (`inclusionCurve()`, `matViz()`, and `plot.covdepGE()`) are included for visualizing the resulting estimates.

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covdepGE-package	<i>covdepGE: Covariate Dependent Graph Estimation</i>
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### Description

A covariate-dependent approach to Gaussian graphical modeling as described in Dasgupta et al. (2022). Employs a novel weighted pseudo-likelihood approach to model the conditional dependence structure of data as a continuous function of an extraneous covariate. The main function, `covdepGE::covdepGE()`, estimates a graphical representation of the conditional dependence structure via a block mean-field variational approximation, while several auxiliary functions (`inclusionCurve()`, `matViz()`, and `plot.covdepGE()`) are included for visualizing the resulting estimates.

### Details

The conditional dependence structure (CDS) of a data matrix with  $p$  variables can be modeled as an undirected graph with  $p$  vertices, where two variables are connected if, and only if, the two variables are dependent given the remaining variables in the data. Gaussian graphical modeling (GGM) seeks to capture the CDS of the data under the assumption that the data are normally distributed. This distributional assumption is convenient for inference, as the CDS is given by the sparsity structure of the precision matrix, where the precision matrix is defined as the inverse covariance matrix of the data.

There is extensive GGM literature and many R packages for GGM, however, all make the restrictive assumption that the precision matrix is homogeneous throughout the data, or that there exists a partition of homogeneous subgroups. `covdepGE` avoids this strong assumption by utilizing information sharing to model the CDS as varying continuously with an extraneous covariate. Intuitively, this implies that observations having similar extraneous covariate values will have similar precision matrices.

To facilitate information sharing while managing complexity, `covdepGE` uses an efficient variational approximation conducted under the novel weighted pseudo-likelihood framework proposed by (1). `covdepGE` further accelerates inference by employing parallelism and executing expensive iterative computations in C++. Additionally, `covdepGE` offers a principled, data-driven approach for hyperparameter specification that only requires the user to input data and extraneous covariates to perform inference. Finally, `covdepGE` offers several wrappers around `ggplot2` for seamless visualization of resulting estimates, such as `matViz`, `inclusionCurve`, and the S3 method `plot.covdepGE`.

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**References**

(1) Sutanoy Dasgupta, Peng Zhao, Prasenjit Ghosh, Debdeep Pati, and Bani Mallick. An approximate Bayesian approach to covariate-dependent graphical modeling. pages 1–59, 2022.

**See Also**

Useful links:

- <https://github.com/JacobHelwig/covdepGE>
- Report bugs at <https://github.com/JacobHelwig/covdepGE/issues>

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covdepGE

*Covariate Dependent Graph Estimation*

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**Description**

Model the conditional dependence structure of  $X$  as a function of  $Z$  as described in (1)

**Usage**

```
covdepGE(  
  X,  
  Z = NULL,  
  hp_method = "hybrid",  
  ssq = NULL,  
  sbsq = NULL,  
  pip = NULL,  
  nssq = 5,  
  nsbsq = 5,  
  npip = 5,  
  ssq_mult = 1.5,  
  ssq_lower = 1e-05,  
  snr_upper = 25,  
  sbsq_lower = 1e-05,  
  pip_lower = 1e-05,  
  pip_upper = NULL,  
)
```

```

tau = NULL,
norm = 2,
center_X = TRUE,
scale_Z = TRUE,
alpha_tol = 1e-05,
max_iter_grid = 10,
max_iter = 100,
edge_threshold = 0.5,
sym_method = "mean",
parallel = FALSE,
num_workers = NULL,
prog_bar = TRUE
)

```

### Arguments

X	$n \times p$ numeric matrix; data matrix. For best results, $n$ should be greater than $p$
Z	NULL OR $n \times q$ numeric matrix; extraneous covariates. If NULL, Z will be treated as constant for all observations, i.e.:  $Z \leftarrow \text{rep}(0, \text{nrow}(X))$ <p>If Z is constant, the estimated graph will be homogeneous throughout the data. NULL by default</p>
hp_method	character in $c(\text{"grid\_search"}, \text{"model\_average"}, \text{"hybrid"})$ ; method for selecting hyperparameters from the the hyperparameter grid. The grid will be generated as the Cartesian product of $\text{ssq}$ , $\text{sbsq}$ , and $\text{pip}$ . Fix $X_j$ , the $j$ -th column of X, as the response; then, the hyperparameters will be selected as follows: <ul style="list-style-type: none"> <li>• If "grid_search", the point in the hyperparameter grid that maximizes the total ELBO summed across all <math>n</math> regressions will be selected</li> <li>• If "model_average", then all posterior quantities will be an average of the variational estimates resulting from the model fit for each point in the hyperparameter grid. The averaging weights for each of the <math>n</math> regressions are the exponentiated ELBO</li> <li>• If "hybrid", then models will be averaged over <math>\text{pip}</math> as in "model_average", with <math>\sigma^2</math> and <math>\sigma_\beta^2</math> chosen for each <math>\pi</math> in <math>\text{pip}</math> by maximizing the total ELBO over the grid defined by the Cartesian product of <math>\text{ssq}</math> and <math>\text{sbsq}</math> as in "grid_search"</li> </ul> "hybrid" by default
ssq	NULL OR numeric vector with positive entries; candidate values of the hyperparameter $\sigma^2$ (prior residual variance). If NULL, $\text{ssq}$ will be generated for each variable $X_j$ fixed as the response as:  $\text{ssq} \leftarrow \text{seq}(\text{ssq\_lower}, \text{ssq\_upper}, \text{length.out} = \text{nssq})$ NULL by default
sbsq	NULL OR numeric vector with positive entries; candidate values of the hyperparameter $\sigma_\beta^2$ (prior slab variance). If NULL, $\text{sbsq}$ will be generated for each variable $X_j$ fixed as the response as:

	<pre>sbsq &lt;- seq(sbsq_lower, sbsq_upper, length.out = nsbsq)</pre>
	NULL by default
pip	<p>NULL OR numeric vector with entries in <math>(0, 1)</math>; candidate values of the hyperparameter <math>\pi</math> (prior inclusion probability). If NULL, pip will be generated for each variable <math>X_j</math> fixed as the response as:</p> <pre>pip &lt;- seq(pip_lower, pi_upper, length.out = npip)</pre>
	NULL by default
nssq	positive integer; number of points to generate for ssq if ssq is NULL. 5 by default
nsbsq	positive integer; number of points to generate for sbsq if sbsq is NULL. 5 by default
npip	positive integer; number of points to generate for pip if pip is NULL. 5 by default
ssq_mult	<p>positive numeric; if ssq is NULL, then for each variable <math>X_j</math> fixed as the response:</p> <pre>ssq_upper &lt;- ssq_mult * stats::var(X_j)</pre> <p>Then, ssq_upper will be the greatest value in ssq for variable <math>X_j</math>. 1.5 by default</p>
ssq_lower	positive numeric; if ssq is NULL, then ssq_lower will be the least value in ssq. $1e-5$ by default
snr_upper	<p>positive numeric; upper bound on the signal-to-noise ratio. If sbsq is NULL, then for each variable <math>X_j</math> fixed as the response:</p> <pre>s2_sum &lt;- sum(apply(X, 2, stats::var)) sbsq_upper &lt;- snr_upper / (pip_upper * s2_sum)</pre> <p>Then, sbsq_upper will be the greatest value in sbsq. 25 by default</p>
sbsq_lower	positive numeric; if sbsq is NULL, then sbsq_lower will be the least value in sbsq. $1e-5$ by default
pip_lower	numeric in $(0, 1)$ ; if pip is NULL, then pip_lower will be the least value in pip. $1e-5$ by default
pip_upper	<p>NULL OR numeric in <math>(0, 1)</math>; if pip is NULL, then pip_upper will be the greatest value in pip. If sbsq is NULL, pip_upper will be used to calculate sbsq_upper. If NULL, pip_upper will be calculated for each variable <math>X_j</math> fixed as the response as:</p> <pre>lasso &lt;- glmnet::cv.glmnet(X, X_j) non0 &lt;- sum(glmnet::coef.glmnet(lasso, s = "lambda.1se")[-1] != 0) non0 &lt;- min(max(non0, 1), p - 1) pip_upper &lt;- non0 / p</pre>
	NULL by default
tau	<p>NULL OR positive numeric OR numeric vector of length <math>n</math> with positive entries; bandwidth parameter. Greater values allow for more information to be shared between observations. Allows for global or observation-specific specification. If NULL, use 2-step KDE methodology as described in (2) to calculate observation-specific bandwidths. NULL by default</p>

norm	numeric in $[1, \infty]$ ; norm to use when calculating weights. Inf results in infinity norm. 2 by default
center_X	logical; if TRUE, center X column-wise to mean 0. TRUE by default
scale_Z	logical; if TRUE, center and scale Z column-wise to mean 0, standard deviation 1 prior to calculating the weights. TRUE by default
alpha_tol	positive numeric; end CAVI when the Frobenius norm of the change in the alpha matrix is within alpha_tol. $1e-5$ by default
max_iter_grid	positive integer; if tolerance criteria has not been met by max_iter_grid iterations during grid search, end CAVI. After grid search has completed, CAVI is performed with the final hyperparameters selected by grid search for at most max_iter iterations. Does not apply to hp_method = "model_average". 10 by default
max_iter	positive integer; if tolerance criteria has not been met by max_iter iterations, end CAVI. 100 by default
edge_threshold	numeric in $(0, 1)$ ; a graph for each observation will be constructed by including an edge between variable $i$ and variable $j$ if, and only if, the $(i, j)$ entry of the symmetrized posterior inclusion probability matrix corresponding to the observation is greater than edge_threshold. 0.5 by default
sym_method	character in c("mean", "max", "min"); to symmetrize the posterior inclusion probability matrix for each observation, the $(i, j)$ and $(j, i)$ entries will be post-processed as sym_method applied to the $(i, j)$ and $(j, i)$ entries. "mean" by default
parallel	logical; if TRUE, hyperparameter selection and CAVI for each of the $p$ variables will be performed in parallel using foreach. Parallel backend may be registered prior to making a call to covdepGE. If no active parallel backend can be detected, then parallel backend will be automatically registered using:  doParallel::registerDoParallel(num_workers)  FALSE by default
num_workers	NULL OR positive integer less than or equal to parallel::detectCores(); argument to doParallel::registerDoParallel if parallel = TRUE and no parallel backend is detected. If NULL, then:  num_workers <- floor(parallel::detectCores() / 2)  NULL by default
prog_bar	logical; if TRUE, then a progress bar will be displayed denoting the number of remaining variables to fix as the response and perform CAVI. If parallel, no progress bar will be displayed. TRUE by default

### Value

Returns object of class covdepGE with the following values:

graphs            list with the following values:

- `graphs`: list of  $n$  numeric matrices of dimension  $p \times p$ ; the  $l$ -th matrix is the adjacency matrix for the  $l$ -th observation
- `unique_graphs`: list; the  $l$ -th element is a list containing the  $l$ -th unique graph and the indices of the observation(s) corresponding to this graph
- `inclusion_probs_sym`: list of  $n$  numeric matrices of dimension  $p \times p$ ; the  $l$ -th matrix is the symmetrized posterior inclusion probability matrix for the  $l$ -th observation
- `inclusion_probs_asym`: list of  $n$  numeric matrices of dimension  $p \times p$ ; the  $l$ -th matrix is the posterior inclusion probability matrix for the  $l$ -th observation prior to symmetrization

`variational_params`

list with the following values:

- `alpha`: list of  $p$  numeric matrices of dimension  $n \times (p - 1)$ ; the  $(i, j)$  entry of the  $k$ -th matrix is the variational approximation to the posterior inclusion probability of the  $j$ -th variable in a weighted regression with variable  $k$  fixed as the response, where the weights are taken with respect to observation  $i$
- `mu`: list of  $p$  numeric matrices of dimension  $n \times (p - 1)$ ; the  $(i, j)$  entry of the  $k$ -th matrix is the variational approximation to the posterior slab mean for the  $j$ -th variable in a weighted regression with variable  $k$  fixed as the response, where the weights are taken with respect to observation  $i$
- `ssq_var`: list of  $p$  numeric matrices of dimension  $n \times (p - 1)$ ; the  $(i, j)$  entry of the  $k$ -th matrix is the variational approximation to the posterior slab variance for the  $j$ -th variable in a weighted regression with variable  $k$  fixed as the response, where the weights are taken with respect to observation  $i$

`hyperparameters`

list of  $p$  lists; the  $j$ -th list has the following values for variable  $j$  fixed as the response:

- `grid`: matrix of candidate hyperparameter values, corresponding ELBO, and iterations to converge
- `final`: the final hyperparameters chosen by grid search and the ELBO and iterations to converge for these hyperparameters

`model_details` list with the following values:

- `elapsed`: amount of time to fit the model
- `n`: number of observations
- `p`: number of variables
- `ELBO`: ELBO summed across all observations and variables. If `hp_method` is "model\_average" or "hybrid", this ELBO is averaged across the hyperparameter grid using the model averaging weights for each variable
- `num_unique`: number of unique graphs
- `grid_size`: number of points in the hyperparameter grid
- `args`: list containing all passed arguments of length 1

`weights` list with the following values:

- **weights:**  $n \times n$  numeric matrix. The  $(i, j)$  entry is the similarity weight of the  $i$ -th observation with respect to the  $j$ -th observation using the  $j$ -th observation's bandwidth
- **bandwidths:** numeric vector of length  $n$ . The  $i$ -th entry is the bandwidth for the  $i$ -th observation

## Overview

Suppose that  $X$  is a  $p$ -dimensional data matrix with  $n$  observations and that  $Z$  is a  $q$ -dimensional extraneous covariate, also with  $n$  observations, where the  $l$ -th observation in  $Z$  is associated with the  $l$ -th observation in  $X$ . Further suppose that the  $l$ -th row of  $X$  follows a  $p$ -dimensional Gaussian distribution with mean 0 and precision matrix  $\Omega(z_l)$ , where  $z_l$  is the  $l$ -th entry of  $Z$  and  $\Omega$  is a continuous function mapping from the space of extraneous covariates to the space of  $p \times p$  non-singular matrices. Then, for the  $l$ -th observation, the  $(j, k)$  entry of  $\Omega(z_l)$  is non-zero if, and only if, variable  $j$  and variable  $k$  are dependent given the remaining variables in  $X$ .

Given data satisfying these assumptions, the covdepGE function employs the algorithm described in (1) to estimate a graphical representation of the structure of  $\Omega$  for each of the observations in  $X$  as a continuous function of  $Z$ . This graph contains an undirected edge between two variables  $X_j$  and  $X_k$  if, and only if,  $X_j$  and  $X_k$  are conditionally dependent given the remaining variables. Core components of this methodology are the weighted pseudo-likelihood framework in which inference is conducted via a block mean-field variational approximation.

## Graph Estimation

Graphs are constructed using a pseudo-likelihood approach by fixing each of the columns  $X_j$  of  $X$  as the response and performing a spike-and-slab regression using the remaining variables  $X_k$  in  $X$  as predictors. To determine if an edge should be added between  $X_j$  and  $X_k$ , the posterior inclusion probability of  $X_k$  in a regression with  $X_j$  fixed as the response ( $PIP_j(X_k)$ ) and vice versa ( $PIP_k(X_j)$ ) are symmetrized according to `sym_method` (e.g., by taking the mean of  $PIP_k(X_j)$  and  $PIP_j(X_k)$ ). If the symmetrized  $PIP$  is greater than `edge_threshold`, an edge will be included between  $X_j$  and  $X_k$ .

To model  $\Omega$  as a function of  $Z$ ,  $n$  weighted spike-and-slab regressions are performed for each variable  $X_j$  fixed as the response. The similarity weights for the  $l$ -th regression are taken with respect to observation  $l$  such that observations having similar values of  $Z$  to  $z_l$  will have larger weights. These similarity weights in conjunction with the pseudo-likelihood framework comprise the weighted pseudo-likelihood approach introduced by (1). Note that model performance is best when  $n > p$ .

## Variational Inference

Spike-and-slab posterior quantities are estimated using a block mean-field variational approximation. Coordinate Ascent Variational Inference (CAVI) is performed for each of the weighted regressions to select the variational parameters that maximize the ELBO. The parameters for each of the regression coefficients are the mean and variance of the slab ( $\mu$  and  $\sigma_{\text{var}}^2$ , respectively) and the probability that the coefficient is non-zero ( $\alpha$ ).  $\mu$  and  $\alpha$  for all coefficients are initialized as 0 and 0.2, respectively.

CAVI for the  $n$  regressions is performed simultaneously for variable  $X_j$  fixed as the response. With each of the  $n$  sets of  $\alpha$  as the rows of an  $n \times (p - 1)$  matrix, the CAVI for variable  $X_j$  is ended for



all  $n$  regressions when the Frobenius norm of the change in the  $\alpha$  matrix is less than `alpha_tol` or after `max_iter` iterations of CAVI have been performed.

Note that since the regressions performed for variable  $X_j$  and  $X_k$  fixed as the response are independent of each other, they may be performed in parallel by setting `parallel = TRUE`. Registering parallel backend with greater than  $p$  workers offers no benefit, since each worker takes on one variable to fix as the response and perform the  $n$  regressions.

### Hyperparameter specification

Each regression requires the specification of 3 hyperparameters:  $\pi$  (the prior probability of inclusion),  $\sigma^2$  (the prior residual variance), and  $\sigma_\beta^2$  (the prior variance of the slab). `covdepGE` offers 3 methods for hyperparameter specification via the `hp_method` argument: `grid_search`, `model_average`, and `hybrid`. Empirically, `grid_search` offers the best sensitivity and `model_average` offers the best specificity, while `hybrid` sits between the other two methods in both metrics.

The hyperparameter candidate grid is generated by taking the Cartesian product between `ssq`, `sbsq`, and `pip` (candidate values for  $\sigma^2$ ,  $\sigma_\beta^2$ , and  $\pi$ , respectively). Each of the methods gives an approach for selecting points from this grid.

In `grid_search`, the point from the grid that produces the model that has the greatest total ELBO is selected, where the total ELBO is calculated by summing the ELBO for each of the  $n$  regressions for a variable  $X_j$  fixed as the response. Thus, all observations use the same set of hyperparameters for the regression on  $X_j$ .

Instead of selecting only one model as in `grid_search`, models are averaged over in `model_average`. With  $X_j$  fixed as the response, the unnormalized weights for each grid point used to perform this averaging is calculated by exponentiating the ELBO for each of the  $n$  regressions. Note that since the ELBO for a given grid point will vary across the  $n$  regressions due to differing similarity weights, each of the  $n$  sets of averaging weights will be unique.

Finally, `hybrid` combines `grid_search` and `model_average`. Fixing  $X_j$  as the response, for each  $\pi$  candidate in `pip`, the point in the grid defined by the Cartesian product of `ssq` and `sbsq` is selected by maximizing the total ELBO summed across the  $n$  regressions. The resulting models for each of the  $\pi$  candidates are then averaged using the exponentiated ELBO for each of the  $n$  regressions as the unnormalized averaging weights.

Note that in the search step of `grid_search` and `hybrid`, CAVI for each of the grid points is performed for at most `max_iter_grid` iterations. A second CAVI is then performed for `max_iter` iterations using the hyperparameters that maximized the total ELBO in the first step. Setting `max_iter_grid` to be less than `max_iter` (as is the default) will result in a more efficient search.

### Candidate grid generation

The candidate grids (`ssq`, `sbsq`, and `pip`) may be passed as arguments, however, by default, these grids are generated automatically. Each of the grids are spaced uniformly between an upper end point and a lower end point. The number of points in each grid is 5 by default. Grids include end points, and the number of points in each grid is controlled by the arguments `nssq`, `nsbsq`, and `npip`. The lower endpoints (`ssq_lower`, `sbsq_lower`, and `pip_lower`) are all  $1e-5$  by default. The upper endpoints are calculated dependent on the variable  $X_j$  fixed as the response.

`ssq_upper` is simply the variance of  $X_j$  times `ssq_mult`. By default, `ssq_mult` is 1.5.

`pip_upper` is calculated by regressing the remaining variables on  $X_j$  using LASSO. The shrinkage hyperparameter for LASSO is chosen to be `lambda.1se`. The number of non-zero coefficients

estimated by LASSO is then divided by  $p - 1$  to calculate `pip_upper`. Note that if the LASSO estimate to the number of non-zero coefficients is 0 or  $p - 1$ , this estimate is changed to 1 or  $p - 2$  (respectively) to ensure that `pip_upper` is greater than 0 and less than 1.

Finally, an upper bound is induced on  $\sigma_\beta^2$  by deriving a rough upper bound for the signal-to-noise ratio that depends on  $\sigma_\beta^2$ . Let  $\sum s_j^2$  be the sum of the sample variances of the columns of the predictors  $X'$ . Under the simplifying assumptions that the expected values of  $X'$  and the spike-and-slab regression coefficients  $\beta$  are 0 and that  $X'$  and  $\beta$  are independent, the variance of the dot product of  $X'$  with  $\beta$  is  $\pi \cdot \sigma^2 \cdot \sigma_\beta^2 \cdot \sum s_j^2$ . Thus, the signal-to-noise ratio under these assumptions is given by  $\pi \cdot \sigma_\beta^2 \cdot \sum s_j^2$ . Replacing  $\pi$  with `pip_upper` and  $\sigma_\beta^2$  with `sbsq_upper` gives an upper bound on the signal-to-noise ratio. Setting this bound equal to `snr_upper` gives an expression for `sbsq_upper`.

### Similarity Weights

The similarity weight for observation  $k$  with respect to observation  $l$  is  $\phi_{\tau_l}(\|z_l - z_k\|)$ . Here,  $\|\cdot\|$  denotes the norm specified by the norm argument,  $z_l$  and  $z_k$  are the values of  $Z$  for the  $l$ -th and  $k$ -th observations,  $\phi_{\tau_l}$  is the univariate Gaussian density with standard deviation  $\tau_l$ , and  $\tau_l$  is the bandwidth for the  $l$ -th observation.

`tau` may be passed as an argument, however, by default, it is estimated using the methodology given in (2). (2) describes a two-step approach for density estimation, where in the first step, an initial estimate is calculated using Silverman's rule of thumb for initializing bandwidth values, and in the second step, the density is refined by updating the bandwidth values. This methodology is used here to estimate the density of  $Z$ , and the updated bandwidths from the second step are used for `tau`.

### References

- (1) Sutanoy Dasgupta, Peng Zhao, Prasenjit Ghosh, Debdeep Pati, and Bani Mallick. An approximate Bayesian approach to covariate-dependent graphical modeling. pages 1–59, 2022.
- (2) Sutanoy Dasgupta, Debdeep Pati, and Anuj Srivastava. A Two-Step Geometric Framework For Density Modeling. *Statistica Sinica*, 30(4):2155–2177, 2020.

### Examples

```
library(ggplot2)

# get the data
set.seed(12)
data <- generateData()
X <- data$X
Z <- data$Z
interval <- data$interval
prec <- data$true_precision

# get overall and within interval sample sizes
n <- nrow(X)
n1 <- sum(interval == 1)
n2 <- sum(interval == 2)
n3 <- sum(interval == 3)

# visualize the distribution of the extraneous covariate
```

```

ggplot(data.frame(Z = Z, interval = as.factor(interval))) +
  geom_histogram(aes(Z, fill = interval), color = "black", bins = n %% 5)

# visualize the true precision matrices in each of the intervals

# interval 1
matViz(prec[[1]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 1, observations 1,...", n1))

# interval 2 (varies continuously with Z)
cat("\nInterval 2, observations ", n1 + 1, ", ..., ", n1 + n2, sep = "")
int2_mats <- prec[interval == 2]
int2_inds <- c(5, n2 %% 2, n2 - 5)
lapply(int2_inds, function(j) matViz(int2_mats[[j]], incl_val = TRUE) +
  ggtitle(paste("True precision matrix, interval 2, observation", j + n1)))

# interval 3
matViz(prec[[length(prec)]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 3, observations ",
    n1 + n2 + 1, ", ..., ", n1 + n2 + n3))

# fit the model and visualize the estimated graphs
(out <- covdepGE(X, Z, nssq = 2, nsbsq = 2, npip = 2))
plot(out)

# visualize the posterior inclusion probabilities for variables (1, 3) and (1, 2)
inclusionCurve(out, 1, 2)
inclusionCurve(out, 1, 3)

```

---

generateData

*Generate Covariate-Dependent Data*


---

## Description

Generate a 1-dimensional extraneous covariate and  $p$ -dimensional Gaussian data with a precision matrix that varies as a continuous function of the extraneous covariate. This data is distributed similar to that used in the simulation study from (1)

## Usage

```
generateData(p = 5, n1 = 60, n2 = 60, n3 = 60, Z = NULL, true_precision = NULL)
```

## Arguments

p	positive integer; number of variables in the data matrix. 5 by default
n1	positive integer; number of observations in the first interval. 60 by default
n2	positive integer; number of observations in the second interval. 60 by default
n3	positive integer; number of observations in the third interval. 60 by default

Z	NULL or numeric vector; extraneous covariate values for each observation. If NULL, Z will be generated from a uniform distribution on each of the intervals
true_precision	NULL OR list of matrices of dimension $p \times p$ ; true precision matrix for each observation. If NULL, the true precision matrices will be generated dependent on Z. NULL by default

### Value

Returns list with the following values:

X	a $(n1 + n2 + n3) \times p$ numeric matrix, where the $i$ -th row is drawn from a $p$ -dimensional Gaussian with mean 0 and precision matrix true_precision[[i]]
Z	a $(n1 + n2 + n3) \times 1$ numeric matrix, where the $i$ -th entry is the extraneous covariate $z_i$ for observation $i$
true_precision	list of $n1 + n2 + n3$ matrices of dimension $p \times p$ ; the $i$ -th matrix is the precision matrix for the $i$ -th observation
interval	vector of length $n1 + n2 + n3$ ; interval assignments for each of the observations, where the $i$ -th entry is the interval assignment for the $i$ -th observation

### Extraneous Covariate

If Z = NULL, then the generation of Z is as follows:

The first  $n1$  observations have  $z_i$  from from a uniform distribution on the interval  $(-3, -1)$  (the first interval).

Observations  $n1 + 1$  to  $n1 + n2$  have  $z_i$  from from a uniform distribution on the interval  $(-1, 1)$  (the second interval).

Observations  $n1 + n2 + 1$  to  $n1 + n2 + n3$  have  $z_i$  from a uniform distribution on the interval  $(1, 3)$  (the third interval).

### Precision Matrices

If true\_precision = NULL, then the generation of the true precision matrices is as follows:

All precision matrices have 2 on the diagonal and 1 in the  $(2, 3)/(3, 2)$  positions.

Observations in the first interval have a 1 in the  $(1, 2)/(1, 2)$  positions, while observations in the third interval have a 1 in the  $(1, 3)/(3, 1)$  positions.

Observations in the second interval have 2 entries that vary as a linear function of their extraneous covariate. Let  $\beta = 1/2$ . Then, the  $(1, 2)/(2, 1)$  positions for the  $i$ -th observation in the second interval are  $\beta \cdot (1 - z_i)$ , while the  $(1, 3)/(3, 1)$  entries are  $\beta \cdot (1 + z_i)$ .

Thus, as  $z_i$  approaches  $-1$  from the right, the associated precision matrix becomes more similar to the matrix for observations in the first interval. Similarly, as  $z_i$  approaches 1 from the left, the matrix becomes more similar to the matrix for observations in the third interval.

**Examples**

```

library(ggplot2)

# get the data
set.seed(12)
data <- generateData()
X <- data$X
Z <- data$Z
interval <- data$interval
prec <- data$true_precision

# get overall and within interval sample sizes
n <- nrow(X)
n1 <- sum(interval == 1)
n2 <- sum(interval == 2)
n3 <- sum(interval == 3)

# visualize the distribution of the extraneous covariate
ggplot(data.frame(Z = Z, interval = as.factor(interval))) +
  geom_histogram(aes(Z, fill = interval), color = "black", bins = n %% 5)

# visualize the true precision matrices in each of the intervals

# interval 1
matViz(prec[[1]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 1, observations 1,...", n1))

# interval 2 (varies continuously with Z)
cat("\nInterval 2, observations ", n1 + 1, ", ..., ", n1 + n2, sep = "")
int2_mats <- prec[interval == 2]
int2_inds <- c(5, n2 %% 2, n2 - 5)
lapply(int2_inds, function(j) matViz(int2_mats[[j]], incl_val = TRUE) +
  ggtitle(paste("True precision matrix, interval 2, observation", j + n1)))

# interval 3
matViz(prec[[length(prec)]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 3, observations ",
    n1 + n2 + 1, ", ..., ", n1 + n2 + n3))

# fit the model and visualize the estimated graphs
(out <- covdepGE(X, Z, nssq = 2, nsbsq = 2, npip = 2))
plot(out)

# visualize the posterior inclusion probabilities for variables (1, 3) and (1, 2)
inclusionCurve(out, 1, 2)
inclusionCurve(out, 1, 3)

```

**Description**

Plot the posterior inclusion probability of an edge between two variables as a function of observation index

**Usage**

```
inclusionCurve(
  out,
  col_idx1,
  col_idx2,
  line_type = "solid",
  line_size = 0.5,
  line_color = "black",
  point_shape = 21,
  point_size = 1.5,
  point_color = "#500000",
  point_fill = "white"
)
```

**Arguments**

<code>out</code>	object of class <code>covdepGE</code> ; return of <code>covdepGE</code> function
<code>col_idx1</code>	integer in $[1, p]$ ; column index of the first variable
<code>col_idx2</code>	integer in $[1, p]$ ; column index of the second variable
<code>line_type</code>	linetype; ggplot2 line type to interpolate the probabilities. "solid" by default
<code>line_size</code>	positive numeric; thickness of the interpolating line. 0.5 by default
<code>line_color</code>	color; color of interpolating line. "black" by default
<code>point_shape</code>	shape; shape of the points denoting observation-specific inclusion probabilities; 21 by default
<code>point_size</code>	positive numeric; size of probability points. 1.5 by default
<code>point_color</code>	color; color of probability points. "#500000" by default
<code>point_fill</code>	color; fill of probability points. Only applies to select shapes. "white" by default

**Value**

Returns ggplot2 visualization of inclusion probability curve

**Examples**

```
library(ggplot2)

# get the data
set.seed(12)
data <- generateData()
X <- data$X
Z <- data$Z
```

```

interval <- data$interval
prec <- data$true_precision

# get overall and within interval sample sizes
n <- nrow(X)
n1 <- sum(interval == 1)
n2 <- sum(interval == 2)
n3 <- sum(interval == 3)

# visualize the distribution of the extraneous covariate
ggplot(data.frame(Z = Z, interval = as.factor(interval))) +
  geom_histogram(aes(Z, fill = interval), color = "black", bins = n %% 5)

# visualize the true precision matrices in each of the intervals

# interval 1
matViz(prec[[1]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 1, observations 1,...,", n1))

# interval 2 (varies continuously with Z)
cat("\nInterval 2, observations ", n1 + 1, ", ..., ", n1 + n2, sep = "")
int2_mats <- prec[interval == 2]
int2_inds <- c(5, n2 %% 2, n2 - 5)
lapply(int2_inds, function(j) matViz(int2_mats[[j]], incl_val = TRUE) +
  ggtitle(paste("True precision matrix, interval 2, observation", j + n1)))

# interval 3
matViz(prec[[length(prec)]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 3, observations ",
    n1 + n2 + 1, ", ..., ", n1 + n2 + n3))

# fit the model and visualize the estimated graphs
(out <- covdepGE(X, Z, nssq = 2, nsbsq = 2, npip = 2))
plot(out)

# visualize the posterior inclusion probabilities for variables (1, 3) and (1, 2)
inclusionCurve(out, 1, 2)
inclusionCurve(out, 1, 3)

```

---

matViz

*Visualize a matrix*


---

## Description

Create a visualization of a matrix

## Usage

```

matViz(
  x,

```

```

color1 = "white",
color2 = "#500000",
grid_color = "black",
incl_val = FALSE,
prec = 2,
font_size = 3,
font_color1 = "black",
font_color2 = "white",
font_thres = mean(x)
)

```

### Arguments

x	matrix; matrix to be visualized
color1	color; color for low entries. "white" by default
color2	color; color for high entries. "#500000" by default
grid_color	color; color of grid lines. "black" by default
incl_val	logical; if TRUE, the value for each entry will be displayed. FALSE by default
prec	positive integer; number of decimal places to round entries to if incl_val is TRUE. 2 by default
font_size	positive numeric; size of font if incl_val is TRUE. 3 by default
font_color1	color; color of font for low entries if incl_val is TRUE. "black" by default
font_color2	color; color of font for high entries if incl_val is TRUE. "white" by default
font_thres	numeric; values less than font_thres will be displayed in font_color1 if incl_val is TRUE. mean(x) by default

### Value

Returns ggplot2 visualization of matrix

### Examples

```

library(ggplot2)

# get the data
set.seed(12)
data <- generateData()
X <- data$X
Z <- data$Z
interval <- data$interval
prec <- data$true_precision

# get overall and within interval sample sizes
n <- nrow(X)
n1 <- sum(interval == 1)
n2 <- sum(interval == 2)
n3 <- sum(interval == 3)

```



```

# visualize the distribution of the extraneous covariate
ggplot(data.frame(Z = Z, interval = as.factor(interval))) +
  geom_histogram(aes(Z, fill = interval), color = "black", bins = n %% 5)

# visualize the true precision matrices in each of the intervals

# interval 1
matViz(prec[[1]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 1, observations 1,...,", n1))

# interval 2 (varies continuously with Z)
cat("\nInterval 2, observations ", n1 + 1, ", ..., ", n1 + n2, sep = "")
int2_mats <- prec[interval == 2]
int2_inds <- c(5, n2 %% 2, n2 - 5)
lapply(int2_inds, function(j) matViz(int2_mats[[j]], incl_val = TRUE) +
  ggtitle(paste("True precision matrix, interval 2, observation", j + n1)))

# interval 3
matViz(prec[[length(prec)]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 3, observations ",
    n1 + n2 + 1, ", ..., ", n1 + n2 + n3))

# fit the model and visualize the estimated graphs
(out <- covdepGE(X, Z, nssq = 2, nsbsq = 2, npip = 2))
plot(out)

# visualize the posterior inclusion probabilities for variables (1, 3) and (1, 2)
inclusionCurve(out, 1, 2)
inclusionCurve(out, 1, 3)

```

---

plot.covdepGE

*Plot the Graphs Estimated by covdepGE*


---

## Description

Create a list of the unique graphs estimated by covdepGE

## Usage

```

## S3 method for class 'covdepGE'
plot(x, graph_colors = NULL, title_sum = TRUE, ...)

```

## Arguments

x	object of class covdepGE; return of covdepGE function
graph_colors	NULL OR vector; the $j$ -th element is the color for the $j$ -th graph. If NULL, all graphs will be colored with "#500000". NULL by default
title_sum	logical; if TRUE the indices of the observations corresponding to the graph will be included in the title. TRUE by default
...	additional arguments will be ignored

**Value**

Returns list of ggplot2 visualizations of unique graphs estimated by covdepGE

**Examples**

```
library(ggplot2)

# get the data
set.seed(12)
data <- generateData()
X <- data$X
Z <- data$Z
interval <- data$interval
prec <- data$true_precision

# get overall and within interval sample sizes
n <- nrow(X)
n1 <- sum(interval == 1)
n2 <- sum(interval == 2)
n3 <- sum(interval == 3)

# visualize the distribution of the extraneous covariate
ggplot(data.frame(Z = Z, interval = as.factor(interval))) +
  geom_histogram(aes(Z, fill = interval), color = "black", bins = n %% 5)

# visualize the true precision matrices in each of the intervals

# interval 1
matViz(prec[[1]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 1, observations 1,...,", n1))

# interval 2 (varies continuously with Z)
cat("\nInterval 2, observations ", n1 + 1, ", ..., ", n1 + n2, sep = "")
int2_mats <- prec[interval == 2]
int2_inds <- c(5, n2 %% 2, n2 - 5)
lapply(int2_inds, function(j) matViz(int2_mats[[j]], incl_val = TRUE) +
  ggtitle(paste("True precision matrix, interval 2, observation", j + n1)))

# interval 3
matViz(prec[[length(prec)]], incl_val = TRUE) +
  ggtitle(paste0("True precision matrix, interval 3, observations ",
    n1 + n2 + 1, ", ..., ", n1 + n2 + n3))

# fit the model and visualize the estimated graphs
(out <- covdepGE(X, Z, nssq = 2, nsbsq = 2, npip = 2))
plot(out)

# visualize the posterior inclusion probabilities for variables (1, 3) and (1, 2)
inclusionCurve(out, 1, 2)
inclusionCurve(out, 1, 3)
```

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